

ECE 520.651
Random Signals Analysis

Final Examination

9:00 AM — 12:00 PM, December 19, 2003.

Name: _____

Read these instructions before starting the examination.

- (i) This is an open-book examination. Use of the Stark and Woods textbook and the Poor textbook (original or photocopy) is permitted. Photocopied material from other books, handwritten/class notes, homework solutions, *etc.* are **not** permitted.
- (ii) Use of electronic calculators is permitted for numeric calculations only.
- (iii) Show all your work clearly and concisely. Points may be deducted for illegible or unclear answers.
- (iv) Write your answers in the space provided. Use the unprinted side of the pages for additional space.
- (v) There are four mandatory questions for a total of 80 points.
- (vi) If you prefer to receive your final grade via e-mail, *explicitly* provide your e-mail address below; you will receive it via the registrar as usual if you leave this space blank.

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Best of luck!

Question No 1	/20 Points
Question No 2	/20 Points
Question No 3	/20 Points
Question No 4	/20 Points

TOTAL	/80 Points
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Question No 1: *Characterization of Random Processes*

- (1a) Let $X[n]$ be a real-valued stationary random sequence with mean $E\{X[n]\} = \mu_X$ and autocorrelation function $E\{X[n+m]X[n]\} = R_{XX}[m]$. If $X[n]$ is the input to a D/A converter, the continuous-time output can be idealized as the *analog* random process $X_a(t)$ with

$$X_a(t) = X[n], \quad \text{for } n \leq t < n+1, \quad \forall t \in \mathbb{R}.$$

- (i) Find the mean function $E\{X_a(t)\} = \mu_{X_a}(t)$ as a function of μ_X .
- (ii) Find the autocorrelation function $E\{X_a(t_1)X_a(t_2)\} = R_{X_a X_a}(t_1, t_2)$ in terms of $R_{XX}[m]$.
- (iii) Is the analog signal $X_a(t)$ wide-sense stationary? Elaborate.

Note: To write your answers compactly, you may wish to use the notation $\lfloor t \rfloor$ to denote the largest integer that is not larger than t .

- (1b) Let $X(t)$ be a real-valued continuous-time Markov random process, with $t \in [0, \infty)$, initial density $f_X(x; 0)$ at time $t = 0$, and conditional pdf

$$f_X(x_2|x_1; t_2, t_1) = \frac{1}{\sqrt{2\pi(t_2 - t_1)}} \exp\left\{-\frac{1}{2} \frac{(x_2 - x_1)^2}{t_2 - t_1}\right\}, \quad \forall t_2 > t_1 \geq 0.$$

- (i) Find $f_X(x; t)$ for an arbitrary time $t > 0$ when $f_X(x; 0) = \delta(x - 1)$.
(ii) Find $f_X(x; t)$ for an arbitrary time $t > 0$ when $f_X(x; 0) = \mathcal{N}(0, 1)$.

Question No 2 *Estimation of Nonrandom Parameters.* Let Y_1 and Y_2 be independent Poisson random variables, each with parameter $\lambda \in \Lambda = (0, \infty)$. Define the parameter θ as

$$\theta = e^{-\lambda}.$$

Consider the problem of estimating θ from $\mathbf{Y} = [Y_1 \ Y_2]^T$.

(2a) Show that $T(\mathbf{Y}) = Y_1 + Y_2$ is a complete sufficient statistic for θ .

(2b) Define an estimate $\hat{\theta}$ by

$$\hat{\theta}(\mathbf{Y}) = \frac{1}{2} [g(Y_1) + g(Y_2)],$$

where $g(\cdot)$ is an indicator function defined as

$$g(y) = \begin{cases} 1 & \text{if } y = 0, \\ 0 & \text{if } y \neq 0. \end{cases}$$

Show that $\hat{\theta}$ is an unbiased estimate of θ .

(2c) Find the MVUE of θ .

(2d) Find the maximum-likelihood estimate of θ from \mathbf{Y} and check if it is unbiased.

(2e) Compute the Cramér-Rao bound on the variance of unbiased estimates of θ .

Hint: See Example 5.7-2 on p280 of Stark and Woods for the pmf of $T(\mathbf{Y})$.

Additional Work Space

Question No 3: *Estimation of Random Variables.*

(3a) Use the *orthogonality principle* to show that the minimum mean-squared error

$$\epsilon^2 = E \left[(X - E[X|Y])^2 \right]$$

in estimating a real-valued random variable X from a real-valued random variable Y may be expressed as

$$\epsilon^2 = E [X (X - E[X|Y])] ,$$

or as

$$\epsilon^2 = E [X^2] - E [(E[X|Y])^2] .$$

Generalize to the case where \mathbf{X} and \mathbf{Y} are real-valued random vectors; *i.e.* show that the MMSE matrix is

$$\begin{aligned} \epsilon^2 &= E \left[(\mathbf{X} - E[\mathbf{X}|\mathbf{Y}]) (\mathbf{X} - E[\mathbf{X}|\mathbf{Y}])^T \right] \\ &= E \left[\mathbf{X} (\mathbf{X} - E[\mathbf{X}|\mathbf{Y}])^T \right] \\ &= E \left[\mathbf{X}\mathbf{X}^T \right] - E \left[E[\mathbf{X}|\mathbf{Y}] E^T[\mathbf{X}|\mathbf{Y}] \right] \end{aligned}$$

(3b) Consider the observation model

$$Y_k = \Theta s_k + N_k, \quad k = 1, 2, \dots,$$

where N_1, N_2, \dots are i.i.d. random variables with common density $\mathcal{N}(0, \sigma^2)$, where Θ , with density $\mathcal{N}(\mu, \nu^2)$, is independent of N_1, N_2, \dots , and s_1, s_2, \dots is a known sequence. Note: Example 9.1-2 on p563 in the Stark and Woods textbook works out the LMMSE estimate for a similar problem.

Let $\hat{\theta}_n$ denote the MMSE estimate of Θ given Y_1, \dots, Y_n .

Find recursions for $\hat{\theta}_n$ and for the minimum mean-squared error $E[(\Theta - \hat{\theta}_n)^2]$ by recasting this as a Kalman filtering problem.

Additional Work Space

Question No 4: Binary Hypothesis Testing. Consider a binary hypothesis testing problem in which a \mathbb{R} -valued observation Y is generated under the null hypothesis, H_0 , according to

$$p_0(y) = \frac{1}{2}e^{-|y|}, \quad y \in \mathbb{R},$$

and under the alternate hypothesis, H_1 , according to

$$p_1(y) = \frac{\theta}{2}e^{-\theta|y|}, \quad y \in \mathbb{R},$$

for some constant $\theta \geq 2$.

- (4a) For $\theta = 2$, find the Bayes decision rule and the minimum Bayes risk for testing H_0 versus H_1 with uniform costs and prior $\pi_0 = \frac{1}{3}$.
- (4b) For $\theta = 2$, find the minimax rule and minimax risk for uniform costs.
- (4c) For $\theta = 2$, find the Neyman-Pearson rule and the corresponding detection probability for false-alarm probability $\alpha \in (0, 1)$.
- (4d) Plot the probability of detection of (4c) as a function of α .
- (4e) If θ were fixed but *unknown*, *i.e.*, if we did not assume $\theta = 2$, how would you revise the Neyman-Pearson rule of (4c)? What does this say about the performance of your (revised) rule for testing H_0 versus H_1 , when H_1 is a *composite* hypothesis?

Additional Work Space